

Difference between multivariate observations at T1 and T2

Description

Compute the differences between multivariate observations (frequency or presence-absence data) forming pairs observed at times T1 and T2. Temporal Beta-diversity Indices (TBI) are computed and tested. TBI are dissimilarity indices that measure beta differentiation through time. They are computed separately between T1 and T2 for each site. The difference between species (or abundances-per-species) losses (*B/den*) and species (or abundances-per-species) gains (*C/den*) can be printed out and tested for significance.

Usage

```
TBI <- function(mat1, mat2, method="%difference", pa.tr=FALSE, nperm=99,
  BCD=TRUE, replace=FALSE, test.BC=TRUE, test.t.perm=FALSE,
  save.BC=FALSE, seed.=NULL, clock=FALSE)
```

Arguments

- `mat1,mat2` Two multivariate community composition or gene frequency data matrices (class `data.frame` or `matrix`) with the same number of rows and columns. The rows must correspond to the same objects (e.g. sites) and the columns to the same variables, e.g. species or alleles.
- `method` One of the following dissimilarity coefficients: {"%difference", "ruzicka", "chord", "hellinger", "log.chord", "sorensen", "jaccard", "ochiai", "euclidean"}. See Details. Names can be abbreviated to a non-ambiguous set of first letters. Default: `method="%difference"`.
- `pa.tr` If `pa.tr=TRUE`, the data are transformed to binary (i.e. presence-absence, or *pa*) form. If `pa.tr=FALSE`, they are not.
- `nperm` Number of permutations for the tests of significance of the temporal beta indices and the permutation test of the B–C difference.
- `BCD` If `BCD=TRUE`, the B and C components of the percentage difference ("%difference") and Ružička ("ruzicka") indices are computed and presented in an output table with three columns: *B/den*, *C/den*, $D=(B+C)/den$, where *den* is the denominator of the index, i.e. $(2A+B+C)$ for the percentage difference index and $(A+B+C)$ for the Ružička index. See **Details** and **Value**.
- If `pa.tr` is `TRUE`, the B and C components are the numbers of species lost or gained, and D is either the Sørensen or the Jaccard dissimilarity. In the BCD output table, column B contains *B/den*, *C/den*, $D=(B+C)/den$, as in the case of the percentage difference and Ružička indices.
- If `BCD=FALSE`, that table is not produced. No table can be computed for indices other than the Ružička and percentage difference or their binary forms.

replace	If <code>replace=FALSE</code> (default value), sampling is done without replacement, producing a regular permutation test. If <code>replace=TRUE</code> , sampling is done with replacement for the test of significance; the method is then bootstrapping.
test.BC	If <code>test.BC=TRUE</code> , the difference between species (or abundances-per-species) gains (C/den) and species (or abundances-per-species) losses (B/den) is tested in a parametric paired t -test computed by function <code>t.test()</code> of <code>{stats}</code> .
test.t.perm	If <code>test.t.perm=TRUE</code> , the difference between species (or abundances-per-species) gains (C/den) and species (or abundances-per-species) losses (B/den) is also tested in a permutational paired t -test computed by function <code>t.paired.perm()</code> .
save.BC	If <code>save.BC=TRUE</code> , the original B and C values are returned in a matrix called \$BC, without division by den as in matrix \$BCD.mat.
seed.=NULL	If <code>seed.=NULL</code> , the random number generator keeps going from the point it had reached in previous calculations. If <code>seed.</code> is an integer value instead of NULL, the random number generator is reset using that value. This allows users to repeat exactly a previous calculation launched with the same value of <code>seed.</code> ; the sequence of generated random numbers will be exactly the same.
clock	If <code>clock=TRUE</code> , the computation time is printed. This option is useful to predict the calculation time when n and $nperm$ are large.

Details

For each object, the function tests the hypothesis (H_0) that the *difference* between T1 and T2 for that object belongs to the same statistical population as the differences displayed by the other objects in the data files. If H_0 is rejected, the object is recognized as exceptionally different from the other objects for its difference between T1 and T2.

To fix ideas, an example in palaeoecology — A researcher is studying ancient and modern diatom communities in sediment cores. If a site displays an exceptional difference between T1 and T2, the researcher is justified to examine the reason for that difference. It could, for example, be caused by a change in land use at that site, which has caused the difference to be larger than at the other sites, compared to the differences caused by climatic change at all sites.

The temporal beta diversity indices available in this function belong to four groups, computed in different ways.

- Method "%difference" computes the percentage difference index, erroneously called the Bray-Curtis index in some software packages; it is the quantitative form of the Sørensen index.

¹ This coefficient was first described as a similarity by Steinhaus in the 1940s and by Motyka et al. (1950), then as a distance by Odum (1950) who called it the *percentage difference*. The Bray & Curtis (1957) paper described a new ordination method; it did not describe, pretended to describe, or used the dissimilarity called the “Bray-Curtis index” in some software packages.

The index used by Bray and Curtis (1957) was a similarity corresponding to $(1 - \text{Whittaker's dissimilarity})$; the latter is also called *Whittaker's (1952) index of association*. The data used by Bray and Curtis in their paper were relative abundances, with row values summing to 1; see the description of their calculations on p. 329 of their paper. The percentage difference (alias Bray-Curtis index) gives a different value than Whittaker's dissimilarity when computed on data expressed as relative abundances. Readers can easily demonstrate the inequality of these two indices for any non-trivial pair of species abundance vectors.

Method "ruzicka" computes the Ružička dissimilarity; this is one of the quantitative coefficients corresponding to the Jaccard dissimilarity for binary data. When these indices are used to compute ordinations by principal coordinate analysis, it is recommended to take the square root of the dissimilarities before the ordination analysis because these indices do not have the property of being Euclidean. However, that precaution is not important here; the results of permutation tests will be the same for these dissimilarities, square-rooted or not. If `pa.tr=TRUE`, either the Sørensen or the Jaccard coefficient are obtained by computing these two coefficients.

- Methods "chord" (chord distance), "hellinger" (Hellinger distance) and "log.chord" (log.chord distance) are obtained by transformation of the species data, as recommended by Legendre & Gallagher (2001), followed by calculation of the Euclidean distance. The transformations are carried out through function `decostand` of the `vegan` package. For the log.chord distance, the data are transformed by $y' = \log(y+1)$ using function `log1p()` of R before calculation of the chord transformation and the Euclidean distance. These three distances have the Euclidean property (Legendre & Legendre 2012, Legendre & De Cáceres 2013). If `pa.tr=TRUE`, the binary distance $\sqrt{2} \cdot \sqrt{1 - \text{Ochiai similarity}}$ is obtained from these three coefficients.

- Methods {"jaccard", "sorensen", "ochiai"} implement the Jaccard, Sørensen and $\sqrt{2} \cdot \sqrt{1 - \text{Ochiai similarity}}$ dissimilarities. For these coefficients, the data are first transformed to presence-absence form (`pa.tr` is given the value `TRUE`), then the dissimilarities are computed using the corresponding quantitative coefficients (Ružička, percentage difference and Hellinger).

- The Euclidean distance is also available in this function. It is not recommended for community composition or allele frequency data. One can compute it for log-transformed abundance data that do not contain zeros, or very few zeros (short gradients).

The temporal beta indices are tested for significance using permutation tests. The hypotheses are the following:

- H_0 : the site under study (e.g. a species assemblage) is not exceptionally different between T1 and T2, compared to other sites that have been observed at the same two times. The differences between T1 and T2 all belong to the same statistical population of differences.
- H_1 : the site under study is exceptionally different between times T1 and T2.

In the decomposition of the Ružička and percentage difference dissimilarities or their presence-absence forms (Jaccard, Sørensen), the components B and C are computed as follows:

- b_j is the part of the abundance of species j that is higher at time 1 than at time 2: $b_j = (y_{1j} - y_{2j})$ if $y_{1j} > y_{2j}$; $b_j = 0$ otherwise. B is the sum of the b_j values for all species in the group of species under study. It is the unscaled sum of **species losses** between time 1 and time 2. In the BCD output table `BCD.mat`, column 1 contains B/den where den is the denominator of the index, i.e. $(2A+B+C)$ for the percentage difference index and $(A+B+C)$ for the Ružička index.

- c_j is the part of the abundance of species j that is higher at time 2 than at time 1: $c_j = (y_{2j} - y_{1j})$ if $y_{2j} > y_{1j}$; $c_j = 0$ otherwise. C is the sum of the c_j values for all species in the group of species under study. It is the unscaled sum of **species gains** between time 1 and time 2. In the BCD output table `BCD.mat`, column 2 contains C/den where den is the denominator of the index, i.e. $(2A+B+C)$ for the percentage difference index and $(A+B+C)$ for the Ružička index.

It is thus incorrect to attribute this coefficient to Bray and Curtis, who never claimed having described a new dissimilarity coefficient. See Legendre & Legendre (2012) and Legendre & De Cáceres (2013) for the formulae of these dissimilarity indices.

Note: The percentage difference should be computed on raw abundances, not on relative abundances.

The original values of B and C for each site, without denominator, are also available in the output table BC.

Warning – In real ecological studies, when the TBI test is applied to data where some sites are highly impoverished due to pollution or other extreme environmental situations, this situation may produce sites with very few species (i.e. very low richness) and no species in common for the T1–T2 comparisons due to sampling variation at these impoverished sites. The TBI dissimilarity will be high and the test may indicate a significant T1–T2 difference if most other sites have higher species richness. This would be a correct statistical outcome for the test. When users of the method identify sites showing significant TBI tests in data, they should check the species richness of these sites at T1 and T2. Interpretation of the test results should be done with caution when high and significant TBI indices are associated with very low richness and no species in common between T1 and T2.

Value

Function TBI returns a list containing the following results:

TBI	The vector of Temporal Beta-diversity Indices (TBI) between observations at times T1 and T2 for each object.
p.TBI	A corresponding vector of p-values. Significant p-values (e.g. $p.TBI \leq 0.05$) indicate exceptional objects for the difference of their species composition.
p.adj	The p-values are corrected for multiple testing using function <code>p.adjust</code> of <code>{stats}</code> . The adjustment is done using <code>method="holm"</code> , which is the default option of the <code>p.adjust</code> function.
BCD.mat	An output table with four columns: <i>B/den</i> , <i>C/den</i> , $D=(B+C)/den$, and <i>Change</i> . The value <i>den</i> is the denominator of the index, i.e. $(2A+B+C)$ for the percentage difference index and $(A+B+C)$ for the Ružička index. The decomposition is such that $D = B/den + C/den$. Columns B and C indicate which of the D values are associated with large B (losses) or large C values (gains), before proceeding to the analysis and interpretation of the D values, using environmental or spatial explanatory variables, through regression or classification tree analysis. When $B > C$, the site has lost species or abundances-per-species between time 1 and time 2; this is indicated by a “–” sign in column <i>Change</i> . On the contrary, if $B < C$, the site has gained species or abundances-per-species between time 1 and time 2; this is indicated by a “+” sign in that column. Sites with equal amounts of losses and gains are marked with a “0”. The <i>B/den</i> and <i>C/den</i> values can be plotted in B–C plots, which are informative about the changes that occurred in the data set between the two surveys under study. If <code>pa.tr</code> is TRUE, the B and C components are the numbers of losses and gains of species, and D is either the Sørensen or the Jaccard dissimilarity. If <code>BCD=FALSE</code> , that table is not produced. No table is (or can be) computed for indices other than the Ružička and percentage difference indices or their binary forms.
BCD.summary	An output table with six columns: $mean(B/den)$; $mean(C/den)$; $mean(D)$; $B/(B+C)$ (which is $mean(B/den)$ divided by $mean(D)$); $C/(B+C)$ (which is $mean(C/den)$ divided by $mean(D)$). These values indicate, over all sites, which of

the processes dominated (loss or gain of species or abundances-per-species) when site compositions changed between time 1 and time 2. *Change* has the same meaning as in table BCD.mat; the sign indicates the direction of the mean change over all sites.

- `t.test_B.C` A matrix giving the results of a paired *t*-test (parametric) of significance of the difference between columns *C/den* and *B/den* of the BCD.mat table. If `test.t.perm=TRUE`, the difference between species gains (*C/den*) and losses (*B/den*) is also tested in a permutational paired *t*-test and the permutational *p*-value is shown in the output table. This result provides an overall test of the direction of change over all sites. It helps confirm the asymmetry between species (or abundances-per-species) gains (*C/den*) and species (or abundances-per-species) losses (*B/den*). A star in column *Change* indicates a significant result of the parametric test at the 0.05 level.
- `BC` An output table with two columns: B and C. In this table, the B and C statistics are not divided by a denominator, contrary to the values *B/den* and *C/den* found in the output table BCD.mat.

References

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- Legendre, P. & L. Legendre. 2012. *Numerical Ecology. 3rd English edition*. Elsevier Science BV, Amsterdam.
- van den Brink, P. J. & C. J. F. ter Braak. 1999. Principal response curves: analysis of time-dependent multivariate responses of biological community to stress. *Environmental Toxicology and Chemistry* 18: 138–148.

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Example

Invertebrate communities subjected to insecticide treatment.

As an example in their paper on Principal Response Curves (PRC), van den Brink & ter Braak (1999) used observations on the abundances of 178 invertebrate species (macroinvertebrates and zooplankton) subjected to treatments in 12 mesocosms by the insecticide chlorpyrifos. The mesocosms were sampled at 11 occasions. The data, available in the {vegan} package, are log-transformed species abundances, $y_{\text{transformed}} = \log_e(10*y+1)$.

The data of survey #4 will be compared to those of survey #11 in this example. Survey #4 was carried out one week after the insecticide treatment, whereas the fauna of the mesocosms was considered by the authors to have fully recovered from the insecticide treatment at survey #11.

```
require(vegan)
data(pyrifos)
```

The mesocosms had originally been attributed at random to the treatments. However, to facilitate presentation of the results, they will be listed here in order of increased insecticide doses: {0, 0, 0, 0, 0.1, 0.1, 0.9, 0.9, 6, 6, 44, 44} µg/L.

```
survey4.order = c(38,39,41,47,37,44,40,46,43,48,42,45)
survey11.order = c(122,123,125,131,121,128,124,130,127,132,126,129)
```

```
# Results using abundance data, percentage difference dissimilarity
res1 <- TBI(pyrifos[survey4.order,], pyrifos[survey11.order,], method="%diff", nperm=999,
BCD=TRUE, test.t.perm=TRUE, clock=TRUE)
```

```
# Results using presence-absence data, Sørensen dissimilarity
res2 <- TBI(pyrifos[survey4.order,], pyrifos[survey11.order,], method="sorensen", nperm=999,
BCD=TRUE, test.t.perm=TRUE, clock=TRUE)
```

```
# Identical results (Sørensen dissimilarity) are obtained with
res3 <- TBI(pyrifos[survey4.order,], pyrifos[survey11.order,], method="%diff", pa.tr=TRUE,
nperm=999, BCD=TRUE, test.t.perm=TRUE, clock=TRUE)
```
